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Sequential RAIM : Theory and application to Civil Aviation Needs

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Biography

Abdelrazak Younes works as a Ph.D. candidate at the Ecole Nationale de l'Aviation Civile (ENAC, France). As part of his Ph.D. thesis at the CNS Research Laboratory of the ENAC, he is now working on the application of sequential algorithms to GNSS integrity monitoring and to GNSS/INS hybridization techniques.

Bacem Bakhache obtained the engineer degree in electrical and electronics from the lebanese university, faculty of engineering, Tripoli, in 1996. He received the D.E.A. (Diplôme d'Etude Approfondie) from the Institut National des Sciences Appliquées (INSA) of Rennes, France. He became a Ph.D. candidate at the University of Technology of Troyes (UTT) in October 1997. His current interests is detection of abrupt changes in signals and systems.

Igor Nikiforov received the Ph.D. in automatic control from the Institute of Control Sciences, Moscow, in 1981. He joined the UTT in 1995, where he is currently Professor. His scientific interest is statistical decision theory.

Abdelahad Benhallam obtained his Ph.D. in communications from the Institut National Polytechnique de Toulouse in 1988. His areas of research include satellite communications, radionavigation and non-stationary signal processing. He is currently in charge of the signal processing department of the CNS Research Laboratory of the ENAC.

Abstract

GPS by itself is unsatisfactory as a sole means of navigation for civil aviation users even with a conventional *Receiver Autonomous Integrity Monitor*- ing (RAIM). The principal limitation of RAIM is its availability. Indeed, even in differential mode (DGPS) or with the Selective Availability turned off, there are periods when the five satellites with sufficiently good geometry required for fault detection are not available. These periods can last several minutes and this is even worse with the fault detection and exclusion function. New sequential algorithms can make up the insufficiency of currently used snapshot methods.

Introduction

Inside a Receiver Autonomous Integrity Monitoring (RAIM), the problem of the availability of the Fault Detection function (FD) and the Fault Detection and Exclusion function (FDE) is strongly related to the geometry of the visible satellites constellation. Due to a poor geometry some faults with a small Signal-to-Noise Ratio (SNR) in residuals can produce a significant impact on the position precision without being detected by conventional snapshot RAIM. Therefore, using algorithms adapted to GPS measurements and typical GPS faults, a sequential RAIM can detect such faults in one of the measured pseudoranges or pseudorange rates.

The aim of this paper is to provide the theoretical basis of the sequential RAIM and to show that the performances in terms of availability that can be achieved by using such algorithms will be better than any snapshot method (under Differential GPS condition or with the Selective Availability turned off).

1 BACKGROUND

1.1 System performance requirements

For civil aviation applications, major problems of the existing systems consist of their lack of continuity and integrity. The Radio Technical Commission for Aeronautics (RTCA) has defined in [RTCA98] the Minimum Operational Performance Standards (MOPS) for airborn equipment using the GPS augmented by a Satellite Based Augmentation Systems (SBAS). In this paper we follow these definitions and requirements in order to compare the proposed FD and FDE algorithms with the well-known snapshot approaches.

1.2 Integrity Monitoring

As stated in [RTCA98] SBAS equipment shall have a FDE capability that utilizes redundant GPS and SBAS ranging measurements to provide independant integrity monitoring. So, RAIM will be used whenever SBAS Integrity is not available but it also can be used as a backup to detect some local anomalies that cannot be accounted for in range corrections supplied by a SBAS during precision approach phases of flight (quoting [G97]).

For a given phase of flight three paramaters can be defined, namely : the **Horizontal Alarm Limit** (HAL), i.e. the radius of a horizontal region centered at the true position which contains the indicated horizontal position with the probability $1 - 10^{-7}$ per flight hour ; the **Vertical Alarm Limit** (VAL) which defines a vertical segment which contains the indicated vertical position with the probability $1 - 10^{-7}$ per flight hour ; the **Vertical Alarm Limit** (VAL) which defines a vertical segment which contains the indicated vertical position with the probability $1 - 10^{-7}$ per flight hour (only for precision approach) ; the **Time-to-Alert** (T_A) which is the delay allowed for FD and FDE algorithms to detect or detect and exclude a faulty satellite. As an example, for a Non-Precision-Approach phase of flight (NPA), we have HAL = 0.3NM and $T_A = 10s$ (There is no VAL for NPA).

Here is RTCA's definitions of the availability of the detection and the exclusion functions :

The detection function is defined to be available when the constellation of satellites provides a geometry for which the missed alert and false alert requirements can be met on all satellites being used for the applicable alert limit and time-to-alert. (...) The exclusion function is defined to be available when the constellation of satellites provides a geometry for wich the FDE algorithm can meet the failed exclusion requirement (...). Therefore, exclusion must occur before the duration of a positioning failure exceeds the time-to-alert, and the detection function as defined above must be available after exclusion (quoting [RTCA98]).

Also, in order to qualify an FD or an FDE algorithms, [RTCA98] specifies the false alarm rate (p_{fa}) to $10^{-5}/h$, the missed alarm rate (p_{ma}) to 10^{-3} and the failed exclusion rate (p_{fe}) to 10^{-3} .

1.3 Notions of Protection Levels

The Horizontal Protection Level_{Fault Detection} (HPL_{FD}) is the radius of a circle in the horizontal plane, with its center being at the true position, which describes the horizontal region of protection for which the missed alert and false alert requirements are met for the chosen set of satellites when autonomous fault detection is used.

The Vertical Protection Level_{Fault Detection} (VPL_{FD}) is half the length of a segment on the vertical axis, with its center being at the true position, which describes the vertical region of protection for which the missed alert and false alert requirements are met for the chosen set of satellites when autonomous fault detection is used.

In other words, given a false alarm rate p_{fa} and a missed detection rate p_{md} , a FD algorithm should be able to detect within the time-to-alert any failure that will cause an error above the HPL_{FD} or the VPL_{FD} .

Following the same idea, given a false alert rate p_{fa} , a false exclusion rate p_{fe} , and a missed alert rate p_{ma} , a FDE algorithm should be able to detect and exclude within the time-to-alert any failure that will cause an error above the HPL_{FDE} or the VPL_{FDE} .

We can also define protection levels that will ensure that the FD algorithm remains able to detect a failure within the time-to-alert after the exclusion of a single satellite ; we call this function the FD^* function. These more stringent corresponding protection levels are noted HPL_{FD}^* and VPL_{FD}^* .

For a given phase of flight, the availability of the FD or the FDE fonctions will be determined by comparing the computed Protection Level(s) with the specified Alarm Limit(s) (c.f. 1.1). For non-

precision approach phases of flight, the different availabilities are defined as :

- FD function available if $HPL_{FD} < HAL$
- FDE function available if $HPL_{FDE} < HAL$
- FD^* function available if $HPL^*_{FD} < HAL$

2 GNSS Navigation and Failure impact

2.1 Position Solution

2.1.1 Notations and definitions

n	Number of Satellites
n	Number of Satellites

- t_j Discretized time
- PR Pseudorange measurement $n \times 1$ vector
- \widehat{PR} Pseudorange estimation $n \times 1$ vector
- Y Linearized measurement $n \times 1$ vector
- $b \sim N(0, W)$ Measurement noise $n \times 1$ vector

$$W = diag(\sigma_1^2, ..., \sigma_n^2)$$
 Measurement noise covari-
ance $n \times n$ matrix

P_{WGS}	True receiver position		
P_{Sat}	Satellites postions		
\widehat{P}_{WGS}	Estimated receiver position		
X	Linearized state 4×1 vector		
Â	Least square estimation of the state 4×1 vector		
w	Least square residuals $n\times 1$ vector		
Н	Direction cosine $n \times 4$ matrix		
$G = (H^T W^{-1} H)^{-1} . H^T W^{-1}$ Projection $4 \times n$ matrix			
P = HG	Idempotent $n \times n$ matrix		
$Q = I_{n \times n}$	$-P$ Residuals projection $n \times n$ matrix		

2.1.2 System linearization

$$Y = PR(t_j) - \widehat{PR}(t_j) \tag{1}$$

$$X = P_{WGS}(t_i) - \widehat{P}_{WGS}(t_i) \tag{2}$$

$$Y = HX + b \tag{3}$$

2.1.3 Least square solution

The retained solution \widehat{X} minimizes the impact of the measurement noise b:

$$\widehat{X} = G.Y = G.H.X + G.b \tag{4}$$

$$w = Q.Y \tag{5}$$

$$\widehat{P}_{WGS}(t_j) = \widehat{P}_{WGS}(t_{j-1}) + \widehat{X}$$
(6)

$$\widehat{PR}(t_j) = \left\| \widehat{P}_{WGS}(t_j) P_{Sat}(t_j) \right\|$$
(7)

It results from these equations that the position error (G.b) is a four-dimensional centered Gaussian variable. It follows by projection of this error on the horizontal plane and on the vertical axis that the horizontal and vertical position errors are centered Gaussian variables which standard deviations are respectively :

$$\sigma_H = \sqrt{\sum_{k=1}^{n} (G_{1k}^2 + G_{2k}^2) . \sigma_k^2}$$
(8)

$$\sigma_V = \sqrt{\sum_{k=1}^n G_{3k}^2 \cdot \sigma_k^2} \tag{9}$$

2.2 Failure Impact

2.2.1 Notations and definitions

Suppose that the current GPS constellation counts n visible satellites at time t and that a failure occurs at time T_F on satellite number k $(1 \le k \le n)$. This failure will be represented as an additional bias B in pseudorange PR_k . We assume that this bias may be positive or negative : we will name type-2k failure a positive failure on satellite k and type-2k + 1 failure a negative failure on satellite k. This convention is not usefull at all for the snaphot RAIM (cf. §3.1) but it is required for the sequential RAIM as we will see in §3.2. Here are the resulting notations :

hypothesis 0 No failure

hypothesis l Type-l failure

- l = 2k Positive failure on satellite k
- l = 2k 1 Negative failure on satellite k
- $(-1)^l B$ Magnitude of type-*l* failure
- $\Gamma_l = (0 \dots 0 (-1)^l B 0 \dots 0)^T$ Resulting measurement error vector
- EH(l) Resulting horizontal position error
- EV(l) Resulting vertical position error
- X_0, Y_0, w_0 Variables under normal conditions (no failure)
- 2.2.2 Least Squares Solution in case of a type-*l* failure

System modelling :

$$Y = HX + b + \Gamma_l \tag{10}$$

$$=Y_0+\Gamma_l \tag{11}$$

Least square solution :

$$w = Q.Y \tag{12}$$

$$= w_0 + Q.\Gamma_l \tag{13}$$

Resulting position error :

$$\hat{X} = G.Y \tag{14}$$

$$= X_0 + G.\Gamma_l \tag{15}$$

The position error $G.\Gamma_l + Gb$ is a four-dimensional Gaussian variable with $G.\Gamma_l$ mean. By projection on the horizontal plane and on the vertical axis, we can deduce the following :

$$EH(l) \sim N\left(\sqrt{G_{1k}^2 + G_{2k}^2} \cdot B, \sigma_H\right) \quad (16)$$

$$EV(l) \sim N (|G_{3k}| B, \sigma_V)$$
(17)

where σ_H and σ_V have been defined in §2.1.3. As these errors do not depend on the sign of the failure, we will note them EH_k and EV_k where k is the faulty satellite in a type-l failure (l = 2k or l = 2k - 1).

3 Integrity Monitoring

3.1 snapshot RAIM

3.1.1 FD fonction

The decision criterion of the detection function is the weighted norm of the least squares residuals vector w:

$$r_{FD} = \frac{\|w\|}{\sqrt{n-4}} \tag{18}$$

Under normal conditions (i.e. no failure), r_{FD}^2 will follow a central χ^2 -distribution with (n-4) degrees of freedom as it is a quadratic sum of centered gaussian variables. The detection threshold h_{FD} is set so that $P(r_{FD} > h_{FD}) = pfd_{FD}$ where pfd_{FD} is the required false detection probability of this algorithm $(pfd_{FD}$ only depends on the required false alert rate p_{fa}).

For a given type-*l* failure of magnitude *B* (affecting satellite *k*), the horizontal position error is a gaussian variable : $EH_k \sim N(\overline{EH_k}, \sigma_H)$ (c.f. §2.2.2); for a given detection threshold, the missed detection rate for a type-*l* failure is $pmd_{FD} = P(r_{FD} < h_{FD})$ where r_{FD}^2 follows a non central χ^2 -distribution with (n-4) degrees of freedom and centrality parameter $\gamma = ||Q\Gamma_l||$. If the measurement noise vector is equal to zero, the value for EH_k which would give $r_{FD} = h_{FD}$ is approximated by ARP_k which is computed as :

$$ARP_k = h_{FD} \times HSlope_k \tag{19}$$

$$HSlope_{k} = \sqrt{\frac{(G_{1k}^{2} + G_{2k}^{2})(n-4)}{Q_{kk}}} \qquad (20)$$

where $HSlope_k$ relates the horizontal position error to the criterion r_{FD} .

So, for a type-*l* failure and without any measurement noise, the HPL_{FD} should be set to ARP_k ; but the existing noise will spread the horizontal positionning impact of this failure around ARP_k . Then, in order to detect this failure with a given probability of detection $(1 - pmd_{FD})$, we shall set HPL_{FD} to $ARP_k + \alpha(pmd_{FD}) \times \sigma_H$. Now, if we consider all the possible failure directions, the horizontal protection level for fault detection is defined by :

$$HPL_{FD} = \max_{1 \le k \le n} \{ARP_k\} + \alpha(pmd_{FD}) \times \sigma_H$$

 $\alpha(p) = \sqrt{2} \operatorname{erf}^{-1}(1-2p)$ is the threshold for which the probability that a variable with a normal distribution will exceed it is equal to p.

Analogously, the vertical protection level for fault detection is defined by :

$$VPL_{FD} = h_{FD} \times VSlope_{Max}$$
 (21)

$$+\alpha(pmd_{FD}) \times \sigma_V$$
 (22)

$$VSlope_{Max} = \max_{1 \le k \le n} \{VSlope(k)\}$$
(23)

$$VSlope(k) = \frac{|G_{3k}|}{\sqrt{Q_{kk}}}$$
(24)

Requirement appliance : If we assume that the noise is uncorrelated between measurements, the false detection probability should be set to the false alert rate per measurement and the missed detection probability should be set to the missed alert rate per measurement :

$$pfd_{FD} = p_{fa} \times \Delta t \tag{25}$$

$$pmd_{FD} = (p_{ma})^{\Delta t/T_A} \tag{26}$$

where Δt is the measurement period.

3.1.2 FDE function

Within a snapshot RAIM, the exclusion algorithm is started only when a failure is detected. In classical architectures, the exclusion function consists in computing the mean square residuals for each subset of n-1 satellites. Then, the subset of satellites which presents the lowest residuals is the one that include no faulty satellite (c.f. [PA88]). But this exclusion function definition does not take into account the notion of probability of false exclusion. In order to satisfy the requirement on false exclusion rate, the fault detection algorithm should be applied to the n subsets of n-1 satellites. Within the exclusion algorithm, the detection threshold should be set so that its false detection probability is equal to the specified p_{fe} (c.f. §1.2). The latter has not been demonstrated and have no theoretical basis but simulations have shown that the resulting false exclusion rate is always under the specified p_{fe} . It results that the protection levels are defined as :

$$HPL_{FDE} = \max\left\{HPL_{FD}^{j}\right\}$$
(27)

$$VPL_{FDE} = \max\left\{VPL_{FD}^{j}\right\}$$
 (28)

where HPL_{FD}^{j} is the HPL_{FD} of the subset excluding satellite number j but which is computed with different parameters than in §3.1.1.

Requirement appliance : If we assume that the FDE function must perform as soon as a failure has been detected, it should require only one set of measurements (as opposed to $T_A/\Delta t$ measurements within the FD function); then the missed detection probability should be set to the missed alert rate. The false detection probability of the FD algorithm applied to each subset of n-1 satellites will be set to the specified false exclusion rate :

$$pfd_{FDE} = p_{fe} \tag{29}$$

$$pmd_{FDE} = p_{ma} \tag{30}$$

3.1.3 FD^* function

The FD^* function is available if all of the FD functions for the *n* subset of n-1 satellites are available. Then the parameter settings are the same as in §3.1.1. The resulting protection levels are definded as :

$$HPL_{FDE} = \max\left\{HPL_{FD}^{j}\right\}$$
(31)

$$VPL_{FDE} = \max\left\{VPL_{FD}^{j}\right\}$$
 (32)

where HPL_{FD}^{j} is the HPL_{FD} of the subset excluding satellite number j (see §3.1.1).

3.2 Sequential RAIM

3.2.1 Introduction

It is known that the sequential algorithms show high performances in detection of stochastic signals and system changes [BN93]. Based on the theory of hypotheses testing, these algorithms would make up the insufficiency of the existing snapshot methods used in GPS integrity monitoring. Instead of testing the least square residuals of the current GPS position and velocity resolution (what is done by commonly used snapshot RAIM), the sequential method accumulates these residuals and detects mean changes occurrence in these residuals. The sequential approach has two advantages over the snapshot approach : the capacity to detect faults with low signal-to-noise ratio in residuals (with a small detection delay); the essentially higher efficiency in the fault isolation step (namely, the lower probability of false isolation). This is a very powerful approach because a drift-type error could be detected before the failure occurs (according to civil aviation specifications). Moreover, these sequential algorithms do have clear theoretical results in terms of missed detection rate, false isolation rate and detection and isolation delays that are very helpful while designing an Integrity Module with respect to the Civil Aviation specifications.

3.2.2 The Log-likelihood ratio

Let us consider a time varying system based on observations $Y_0,...,Y_t$. If the probability density function of these observations is equal to f_1 , then this system is said to satisfy hypothesis H_1 , if it is equal to f_2 , then the system is said to satisfy hypothesis H_2 . The log-likelihood ratio between H_1 and H_2 is a function of time t and is defined as $S_t(H_1, H_2) =$ $\log \frac{f_1(Y_0,...,Y_t)}{f_2(Y_0,...,Y_t)}$. As a consequence, if H_1 is more likely than H_2 $(f_1(Y_0,...,Y_t) > f_2(Y_0,...,Y_t))$, the log-likelyhood ratio $S(H_1, H_2)$ is positive and will increase with time.

3.2.3 Detection and exclusion rules

In the case of a single satellite failure, for a known failure magnitude without a priori assumption on the direction of the failure, the detection and exclusion algorithm will test each possible direction of the failure (c.f. [N95]). If there are n visible satellites, there will be 2n possible directions because the change could be either positive or negative for each satellite. Hence, there are a total of 2n + 1 hypotheses, including hypothesis 0 (c.f. §2.2.1). The sequential theory allows the recursive calculation of the log-likelihood ratio between different hypotheses.

Log-likelihood ratio at time t between hypotheses land 0 (l = 2k or l = 2k - 1) given $S_0(l, 0) = 0$:

$$S_t(l,0) = \left\{ S_{t-1}(l,0) + \frac{(-1)^l \tilde{v} Y(k)}{\sigma_k^2 \sqrt{Q_{kk}}} - \frac{\tilde{v}^2}{2\sigma_k^2} \right\}^+$$

where $\{x\}^+ = \max\{x, 0\}$ and \tilde{v} is a tuning parameter (Idealy $\tilde{v} = B\sqrt{Q_{kk}}$ as stated in §3.2.4).

Log-likelihood ratio between hypotheses l et q at time t:

 $S_t(l,q) = S_t(l,0) - S_t(q,0)$ (33)

where l and q are elements of 1, ..., 2n.

Intuitively, one might say that at least one loglikelihood ratio $S_t(l, 0)$ is positive if there is a failure and that hypothesis l is correct if all of the loglikelihood ratios $S_t(l, q)$ between hypotheses l and q are positive (c.f. §3.2.2). If the hypothesis l remains true, all $S_t(l,q)$ will increase with time because this hypothesis will be more likely than the others. If h_D and h_E are the detection and exclusion thresholds, the decision rules for detection and detection-exclusion are :

$$r_{FD}(t) = \max_{1 \le l \le 2n} \left[S_t(l,0) - h_D \right]$$
(34)

$$r_{FDE}^{l}(t) = \min_{0 \le q \ne l \le 2n} \left[S_t(l,q) - h_q \right]$$
(35)

where $h_q = h_D$ if q = 0 and $h_q = h_E$ if $q \neq 0$. It results that the stopping times for detection and for detection and exclusion are defined as :

$$T_{FD} = \min \{ t \ge 1 : r_{FD}(t) > 0 \}$$
(36)

$$T_{FDE} = \min\left\{t \ge 1 : \max_{1 \le l \le 2n} \left[r_{FDE}^{l}(t)\right] > 0\right\} 37)$$

In the case of exclusion, the excluded satellite is the one for which the exclusion threshold is reached first. In other words, satellite k is excluded if $r_{FDE}^{2k}(t) > 0$ or $r_{FDE}^{2k-1}(t) > 0$.

3.2.4 Parameter Settings

The tuning parameter \tilde{v} : For an a priori known failure magnitude B in the pseudorange of satellite number k, this algorithm is optimal only if parameter \tilde{v} is equal to $B\sqrt{Q_{kk}}$. Because of the regression model, the optimal value for \tilde{v} will vary from satellite to satellite and should be set in accordance. But in general case, the magnitude failure Bis also unknown. To solve this problem, we propose to apply several CUSUM tests in parallel in order to cover a large range of magnitudes $[\tilde{v}_{min}, \tilde{v}_{max}]$. So, the fault detection algorithm will be composed of L parallel CUSUMs with parameters $\tilde{v}_1, \tilde{v}_2, ... \tilde{v}_L$ ranging from \tilde{v}_{min} to \tilde{v}_{max} . As was shown in [N98], the choice of L and parameters \tilde{v}_i can be made so as to minimize the asymptotic theoretical value of the detection delay, L is the smallest integer that verify :

$$L \ge \ln \frac{v_{max}}{v_{min}} \left\{ \ln \frac{e+1}{e-1} \right\}^{-1}$$
(38)

$$v_i = v_{min} \frac{(e+1)^i}{e(e-1)^{i-1}}$$
(39)

where e is defined as the asymptotic efficiency of the CUSUM (the more e is close to 1, the greater L will be). It shall be noticed that these parameters are not designed to minimize the asymptotic theoretical value of the exclusion delay; but montecarlo simulations have shown that the resulting simulated mean isolation delay follows its theoretical value (c.f. [YNB97]).

The thresholds : According to sequential theory (c.f. [W47]), the detection threshold in a CUSUM should be set to $\sim \ln (n/pfd_{FD})$ as $pfd_{FD} \rightarrow 0$ so as to get a given false detection probability pfd_{FD} . If we use L parallel CUSUMs, then the detection thresholds become :

$$h_D \sim \ln\left(\frac{n \times L}{pfd_{FD}}\right)$$
 (40)

Unfortunately, there is no theoretical result concerning the threshold for exclusion but Monte-Carlo simulations results show that, if we set $h_{FDE} \times$ ln (L/pfe_{FDE}) , where pfe_{FDE} is the required false exclusion probability, the resulting false exclusion probability is well under the required one ; but we have to emphasize that this value have no theoretical basis and might lead to wrong results.

3.2.5 Distribution law of the delay of detection τd_D

It was shown by Wald that the detection delay for a type-*l* failure τ_D^l follows (asymptotically) a normal distribution $N(\overline{\tau}_D^l, \sigma_D^l)$ (c.f. [W47]); after simplifications we get :

$$\overline{\tau}_D^l = \frac{h_D}{K_D^l} , \ \sigma_D^l = \frac{\sqrt{2h_D}}{K_D^l}$$
(41)

where $K_D^l = \frac{B^2}{2\sigma_k^2} Q_{kk}$ is the Kullback information relative to hypothesis l.

3.2.6 Distribution law of the delay of detection-exclusion τ_E

Monte-Carlo simulations show that the detection and exclusion delay of a type-*l* failure (τ_E^l) follows a normal distribution $N(\overline{\tau}_E^l, \sigma_E^l)$. The only clear theoretical result concerning exclusion is on $\overline{\tau}_E^l$ (c.f. [W47]). The expression below for σ_E^l has been obtained by a generalisation of the detection problem to the detection-exclusion problem. But simulations have shown that it is a good estimation of the standard deviation of variable τ_E^l :

$$\overline{\tau}_E^l = \max(\overline{\tau}_D^l, \frac{h_E}{K_E^l}), \, \sigma_E^l = \max(\sigma_D^l, \frac{\sqrt{2h_E}}{Z_E^l}) \quad (42)$$

where :

$$K_E^l = B^2 Q_{kk} \min_{j \neq k} \left\{ \frac{\beta_{kj}}{\sigma_{kj}^2} \right\}$$
(43)

$$\sigma_{kj} = \sqrt{2} \left(\frac{1}{\sigma_k^2} + \frac{1}{\sigma_j^2} \right)^{-\frac{1}{2}}$$
(44)

$$\beta_{kj} = \min\left[1 \pm \frac{Q_{kj}}{\sqrt{Q_{kk} + Q_{jj}}}\right]$$
(45)

3.2.7 FD function

We have to ensure that, for a given type-*l* failure, the detection delay τ_D^l will be lower than the specified time-to-alert T_A (c.f. §1.2) with a suitable probability 1 - pmd, where pmd is the required missed detection probability. This consideration brings the following condition :

$$\overline{\tau}_E^l + \alpha(pmd) \times \sigma_E^l \le T_A \tag{46}$$

Combining these equations we get the expression of the minimum detectable magnitude of a type-*l* failure (i.e. $B \ge B_k^{md}$):

$$B_k^{md} = \sigma_k \sqrt{\frac{2}{T_A Q_{kk}} \left(h_D + \alpha(pmd).\sqrt{2h_D}\right)} \quad (47)$$

The horizontal and vertical components of this failure magnitude leads to the expressions of the protection levels for detection :

$$HPL_{FD} = \max_{k=1..n} \left\{ B_k^{md} \sqrt{G_{1k}^2 + G_{2k}^2} \right\}$$
(48)

$$VPL_{FD} = \max_{k=1..n} \left\{ B_k^{md} |G_{3k}| \right\}$$
(49)

3.2.8 FDE function

We have to ensure that, for a given type-*l* failure, the exclusion delay τ_D^l will be lower than the specified time-to-alert T_A (c.f. §1.2) with a suitable probability 1 - pmd, where pmd is the missed detection probability. This consideration brings the following condition :

$$\overline{\tau}_I^l + \alpha(pmd) \times \sigma_I^l \le T_A \tag{50}$$

Combining these equations the minimum excludable magnitude of a type-*l* failure (i.e. $B \ge B_k^{me}$) is given by :

$$(B_k^{me})^2 = \frac{1}{T_A Q_{kk}} \max \left[2\sigma_k^2 h_D, \max_{j \neq k} \left(\frac{\sigma_{kj}^2}{\beta_{kj}} \right) . h_E \right]$$

+ $\frac{\alpha(pmd)}{T_A Q_{kk}} \max \left[2\sigma_k^2 \sqrt{2h_D}, \max_{j \neq k} \left(\frac{\sigma_{kj}^2}{\beta_{kj}} \right) . \sqrt{2h_E} \right]$

The horizontal and vertical components of this failure magnitude leads to the expressions of the horizontal and vertical protection levels for detection and exclusion :

$$HPL_{FE} = \max_{l=1..2n+1} \left\{ B_k^{me} \sqrt{G_{1k}^2 + G_{2k}^2} \right\}$$
(51)

$$VPL_{FE} = \max_{l=1..2n+1} \{B_k^{me} G_{3k}\}$$
(52)

3.2.9 FD^* function

As in §3.1.3, the FD^* function is available if all of the FD functions for the *n* subset of n-1 satellites are available. Then the parameter settings are the same as in §3.2.7. The resulting protection levels are definded as :

$$HPL_{FDE} = \max\left\{HPL_{FD}^{j}\right\}$$
(53)

$$VPL_{FDE} = \max\left\{VPL_{FD}^{j}\right\}$$
 (54)

where HPL_{FD}^{j} is the HPL_{FD} of the subset excluding satellite number j (see §3.2.7).

3.2.10 Requirement appliance

The specified false alert rate and false exclusion rate must be transposed on a per measurements basis, but the missed alert rate has to be considered on a time-to-alert basis :

$$pfd_{FD} = p_{fa} \times \Delta t \tag{55}$$

$$pfe_{FDE} = p_{fe} \times \frac{\Delta t}{T_A}$$
 (56)

$$pmd = p_{ma} \tag{57}$$

4 snapshot RAIM versus sequential RAIM

4.1 Hypotheses for simulation

GPS As it has been annonced that the *Selective* Availability (S.A.) will be turned off in a few years, the chosen error model is based on Standard Positionning System (SPS) without S.A. characteristics. The applied pseudorange measurement noise is then a normally distributed noise with a 12.5m standard

deviation. The GPS pseudoranges used in the simulations are generated by a GPS constellation simulator based on an almanac file (GPS week 973). At this time, 27 satellites were available. The degraded constellations are obtained by disabling the first satellite(s).

Geostationary Satellites (Geo) A GNSS system may include one or more additional geostationary satellites ; the INMARSAT satellites that will be used in future SBAS systems are named AOR-E (Atlantic ocean region east), AOR-W (Atlantic ocean region west), POR (Pacific ocean region) and IOR (Indian ocean region).

4.2 Criteria for comparison

We will use three criteria for comparison of the two RAIM methods, that is to say the HPL_{FD} , the HPL_{FDE} and the HPL_{FD}^* . The considered phase of flight is the Non-Precision approach (NPA) which specifies HAL = 555m and $T_A = 10s$ (c.f. §1.1).

4.3 Simulation results

The plots in figures 1, 2 and 3 represent the evolution of the HPL_{FD} , the HPL_{FDE} and the HPL_{FD}^* as fonctions of time for both sequential (dotted) and snapshot (solid) algorithms. Time ranges from 1998-09-05 0h00 to 1998-09-06 0h00 with a 2 minutes step. GNSS receiver position is located in Toulouse (France).

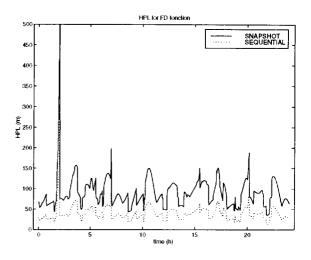
In table 1, the FD, FDE and FD^* functions availabilities are given for Snapshot RAIM (left value) and Sequential RAIM (right values) in regards to NPA requirements. These values are the percentage of time when $HPL \leq HAL$.

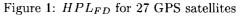
Conclusion

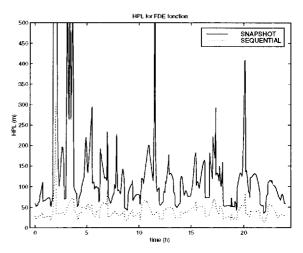
The key problem concerning GPS is the RAIM function availability. HPL_{FD} ploting on figure 3.1.1 shows that there is an improvement in fault detection availability by using Sequential RAIM rather than Snapshot RAIM, and the improvement is very substantial with regards to fault detection and exclusion (see fig. 2). This is even more substantial when the FD has to remain available after a satellite exclusion (see fig. 3). These results show that the constellation geometry affects much less the Sequential RAIM than the snapshot RAIM and that the Sequential exclusion function is much more powerful. These improvements, as summarized in table 1, lead to the conclusion that a SBAS system should use of a Sequential RAIM in order to satisfy NPA requirements for example. As for precision approach where RAIM function needs to be used as a backup, Sequential RAIM will surely outperform Snapshot RAIM in detection of local small errors.

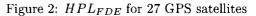
Snap. / Seq.	FD (%)	FDE (%)	FD^{*} (%)
27GPS+3Geo	100/100	100/100	100/100
27GPS+1Geo	100/100	100/100	95.42/100
27GPS	99.86/100	98.06/100	72.82/98.20
26GPS	99.86/100	96.95/100	64.91/97.64
25GPS	99.86/100	92.09/100	51.04/93.20

Table 1:NPARAIMavailability(snapshot/sequential)









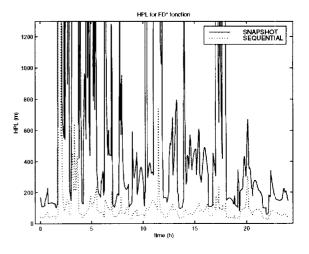


Figure 3: HPL_{FD}^* for 27 GPS satellites

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